



Evaluation of single-arm studies of oncothermia



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Objective

Oncotherapy survival studies are problematic due to the missing control arm. This is a problem in general, when the treatment targets advanced, mostly refractory, relapsed malignancies in high treatment lines, when the only way is the sequential treatment. Usually the care in high line treatment (or in terminal phase) has very limited evidence based possibilities, so medical decision making processes are usually well tailored to the individual patients [18], [19]. In these cases evidences have to be shown when randomized controlled trials are not possible. [20] The sequential trial [1], [2], [3], is well known, and applied frequently in the case of small trials [4]. The sequential trial (like the oncotherapy) is applied for the same patient in sequences, in this approach the development of the patient is measured and documented. Our objective is to show how the evaluation of the single arm study could be realistic enough to be evidence based.

Method

The basic of the idea of the data separation is the appropriate parameterization of the non parametric Kaplan Meier survival pattern by poly Weibull fit. However we have some qualitative assumptions:

- Patient starts the new sequence when the previous had not (or had not satisfactory) result
- The new sequence does not add additional (new) information of the patient's state to the actual therapy in the actual sequence
- The effect of the new sequence affects the survival curve, so the studied Kaplan Meier plot includes the information
- The sequence is medically controlled at least on the same way, as were done in previous therapies. No uncontrolled "side therapies" are in use

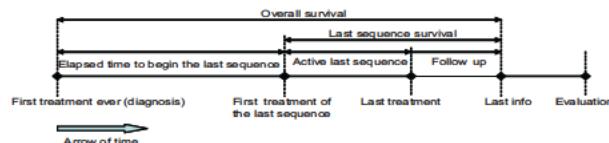
Studying the median of a survival curve alone disregards the real success measurable at the end of the study. Connecting this "mistake" the average (mean) of the distribution is considered. The mean is more affected by the "tail" of the distribution, so it gives more accurate idea on the cure rate. The median is more responsible for the information how quick the loss of the patients, while the mean has more part on the information how long the effect of the high success patients. Both are important for characterization, the time scale and the shape of the distribution are independent parameters. The distribution curve must be characterized at least by two parameters. These two parameters are the mean and the median, supposing to characterize the non parametric distribution, and so in fact this is a hidden parameterization of the Kaplan Meier plot. Best mining of the data would be when the non parametric Kaplan Meier survival plot could be parameterized. Description of survival curves by parametric distribution function was a long time effort, could be approached by fitting the parametric Weibull (Avram) curves [21], [22], [23], [24], [25], on the actual probability function. The universal applicability of the Avram function was recognized much earlier, [26], [27], [28]. Using the Weibull distribution function to approach the survival curve is theoretically and practically established for clinical applications, [25]. Fitting the measured Kaplan Meier survival curve (KM(t)) by a function $S(t)$ composed by two Weibull functions (with parameters denoted by superscripts (R) and (NR)), describing the responders and non responders by a composite ratio C, respectively. Application the parametric Weibull distribution function approaching the survival curve for clinical applications is established theoretically and practically, [5], [6], [7], [8]. It is used for a long time for survival description in gerontology [9], [10] and in oncology [11] as well.

$$KM(t) \approx S(t) = (1-C)e^{-\left(\frac{t}{t_{0(R)}}\right)^{\alpha(R)}} + Ce^{-\left(\frac{t}{t_{0(NR)}}\right)^{\alpha(NR)}}$$

Sequenced trial is well known and applied frequently in the case of small trials [29].

However we have some qualitative assumptions:

1. Patients start the new sequence when the previous had not (or had not satisfactory) result. This condition is generally valid, no reason start new therapy when it works satisfactory. (In some cases due to psychology or other factors anyway could be abandoned a successful therapy but we assume it is less than 5% of all the treatments)
2. The effect of the new sequence affects the survival curve, so the studied Kaplan Meier plot includes the information (Example: when the effect is improving the quality of life but does not affect the survival, the sequence cannot be studied by survival curves)
3. The sequence is medically controlled at least on the same way as were done in previous therapies. No uncontrolled "side therapies" are in use



The time sequences of oncotherapy studies. The time between the first ever treatment and the first oncotherapy is complex, having numerous pre-treatments. It is regarded here as one step.

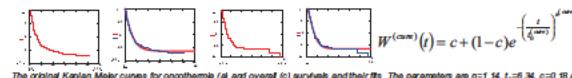
Results

The evaluation of the Kaplan Meier plot by parametric distribution works well in the practice. Patients responding to the treatment are well distinguishable from the nonresponders, and on this basis the overall survival benefit can be evaluated. The significance level depends on the number of patients, but over 25 patients it usually fits better than 95% confidence. We evaluated numerous single arm clinical trials, showing the efficacy of the study. The evaluation was well correlated with the independently measured other parameters as well as the criteria of start of oncotherapy also shows stable reference distribution

$$S^{(hist)}(t) = (1 - C^{(hist)}) \exp\left(-\left(\frac{t}{t_0^{(hist)}}\right)^{\alpha^{(hist)}}\right) + C^{(hist)} \exp\left(-\left(\frac{t}{t_0^{(hist)}}\right)^{\alpha^{(hist)}}\right)$$

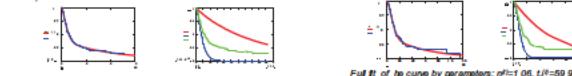
Further control could be given by studying the historical control of the pancreas cancer treatment by the same investigator (n=34), who did the same treatment as the oncotherapy. The results are shown in figure 1. The overall survival curve does not pass fail to detect any significant differences in decomposition. It is a cohort comparison of non responders in overall survival and the control group shows remarkable correspondence this supports again the validity of the decomposition

Let us study an actual example of the pancreas trial (n=99), [12]



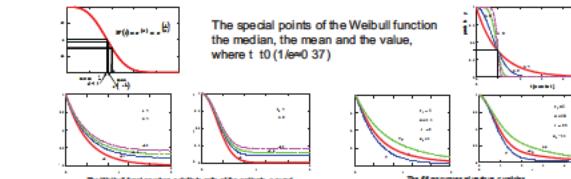
The original Kaplan Meier curves for oncotherapy (a) and overall (c) survivals and their fits. The parameters are n=14, t₀=6.34, c=0.18 and n=138, t₀=13.74, c=0.17 for oncotherapy and overall survival, respectively

Better fits could be achieved by parametric decomposition of the survivals. The decomposition significantly divides the cohort of advanced inoperable pancreatic cancer patients on two subgroups (responders and non responders) in oncotherapy survival. Keeping the C composite parameter the fit and decomposition of the overall survival is available

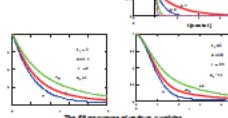


Fit of (a) of the curve by parameters: n=134, n=95, t₀=43.13, t₀=13.74, C=0.75, C=0.61. The decomposition curves show the significant difference of responders (39%) and non responders (61%). Applied the same ratio as in oncotherapy survival

The "inclusion criteria" for the patients to oncotherapy treatment is when the "gold standard" are not eligible. These criteria could be checked by study he elapsed time to the first oncotherapy from the first diagnosis. The time from the first diagnosis to the first oncotherapy has to be a cohort (when the inclusion of the patients to oncotherapy had identical criteria) consequently it has to be characterized by Weibull parametric formulation, where the two distributions are close, or C^{hist}/C^{onco} is small. Indeed fit and decomposes the curve of elapsed time from the first diagnosis to the start of oncotherapy. The definite commands of one single curve. This shows our "inclusion criteria" is really valid cohort forming condition



The special points of the Weibull function
the median, the mean and the value,
where t 10 (1/e=0.37)



The fit of curves of various c values

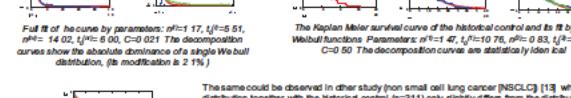
By these assumptions we study a shift of the original cohort distribution fitting it to two groups: responding and non responding patients. The Weibull distribution (10) is divided into two different distributions [11], compared directly, one whom the bias merit had no or minor influence and one where the treatment was effective. The weighted addition of the curves reconstructs the original. The "inclusion criteria" for the patients to oncotherapy treatment is when the "gold standard" are not eligible. These criteria could be checked by study the elapsed time to the first oncotherapy from the first diagnosis. The time from the first diagnosis to the first oncotherapy has to be a cohort (when the inclusion of the patients to oncotherapy had identical criteria) consequently it has to be characterized by single Weibull parametric formulation. The probability density function of the Weibull distribution is considered to be the best fit: the two Weibull curves t₀ and t₀' for each and their composite C (ratio of the non responders) are fitted the patients by their numbers. Introducing the value of the survival at the maximal survival the $S^{(hist)}_{max}$ will be automatically obtained from this fit which is nothing else only the value of the survival at the maximal survival. The $S^{(hist)}_{max}$

The same composite fit is applied for overall survival

$$S^{(per)}(t) = (1-C)\exp\left(-\left(\frac{t}{t_0^{(per)}}\right)^{\alpha^{(per)}}\right) + C\exp\left(-\left(\frac{t}{t_0^{(per)}}\right)^{\alpha^{(per)}}\right) \quad S^{(NR)}(t) = (1-C)\exp\left(-\left(\frac{t}{t_0^{(NR)}}\right)^{\alpha^{(NR)}}\right) + C\exp\left(-\left(\frac{t}{t_0^{(NR)}}\right)^{\alpha^{(NR)}}\right)$$

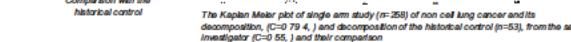
Results

The Kaplan Meier survival curve of the historical control and its fit to two Weibull functions. Parameters: n=117, t₀=10.76, $\alpha=0.83$, $t_0'=10.50$. The decomposition curves are statistically identical



The Kaplan Meier survival curve of the historical control and its fit to two Weibull functions. Parameters: n=117, t₀=10.76, $\alpha=0.83$, $t_0'=10.50$. The decomposition curves are statistically identical

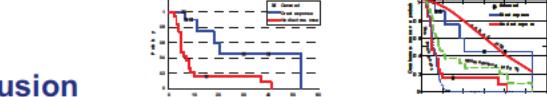
The same could be observed in other study (non small cell lung cancer [NSCLC]) [13], where the distribution together with the historical control (n=311) only slightly differs from the distribution of non responders without the control group (n=29).



The Kaplan Meier plot of single arm study (n=258) of non small cell lung cancer and its decomposition, C=0.79 4, and decomposition of the historical control (n=53), from the same investigator (C=0.65), and their comparison

Other prospective study [14] had measured the local clinical response and the survival time in the same trial. The direct response (CR+PR) shows good, significant correspondence with the parametric separation

Significant correspondence of the measured and calculated separation of the patient's survivals by their local response



Conclusion

Parametric evaluation of Kaplan Meier non parametric distribution works well for single arm studies. The single parameter reference of the Kaplan Meier (median survival) is unsatisfactory; the two parametric Weibull distributions describe the situation much more exactly

References

- [1] Bajusz E, J. Bartha and C. Cheung. 1996. The group sequential integral test for phase II cancer clinical trials. American Journal of Clinical Oncology 9:422-430
- [2] Williams PL. (1998) Sequential monitoring of clinical trials w/ multiple survival endpoints. Statistica Medica 17:2341-2357
- [3] Whittaker J, Donnan J, Donnan AN, Stephen R, Martin D. (1983) A simplified sequential design analysis based on data from two MRC trials. British Journal of Cancer 48:1171-1178
- [4] Evans CH, JR. Statistical SE Trials@2001. Committee on Strategies for Small Number. Participant Clinical Research Trials. Small Clinical Trials: Issues and Challenges. Natl Academies Press. Washington DC. <http://www.nap.edu/catalog/0309073110.html>
- [5] Halpin TA, KO, Harley JA, Joseph L, Collet JP. A Comparison of Parametric and Nonparametric Approaches to ROC Analysis of Organ Specific Diagnostic Tests. Medical Decision Making 10:103-108
- [6] Jones DR, Stukel DA. Multivariate survival analysis with doubly censored data: application to the assessment of Acute care treatment. / Gerontology as a Geriatric Prognostic. Statistics in Medicine 21:2847-2862
- [7] Avram MA. Kinetics of phase change. I. J. Clin. Oncol. 7: 1933-1939
- [8] Avram MA. Kinetics of phase change. II. Mechanism. Mech Ageing Dev 74:15-33 1994
- [9] Plantamour L. A mathematical model of survival kinetics. I. Theoretical basis. Ann Génétol Genér 5:107-118 1998
- [10] Economos AC. Rate scaling as a measure of mortality. Ann Genet Genér 13:27-1982
- [11] Williams CL, Douglas C, Ries L, Lewis D, Smith D, on behalf of UICC-SDG. (2004) Statist. Journal of Oncol 22:225-232
- [12] Szasz A. (2004) Electro hypertermia for advanced pancreatic cancer. Deutscher Kongress für Radiologie e. Strahlentherapie und Mediцинische Physik. Erfurt. 10 Jun.
- [13] Szasz A, Vaskinya A, Mayer T, Szasz O. (2010) Clinical study for advanced breast carcinoma. Biologische Medizin 3:528-530
- [14] Szasz R. (2002) Chemotherapy in cancer: an application to data from the Children's Cancer Group. Statistics in Medicine 21:2933-2942
- [15] Szasz R. (2002) Cure model analysis in cancer: an application to data from the Children's Cancer Group. Journal of Clinical Oncology 20:207-207
- [16] Szasz R. (2007) Optimal Design of Prognostic Risk Groups in Pediatric Cancer: Analysis of Data From the Children's Oncology Group. Journal of Clinical Oncology 25:2070-2077
- [17] Renfro P, Longmate J. (2003) Parametric models for accelerated and proportional survival: a comment on proportional hazards. Statistics in Medicine 21:3279-3289
- [18] Szasz R. (2006) Process and context of Decision making by advanced cancer patients. J Clin Oncol 24:1029-1030
- [19] Tsoarit D, Mostert L, Matot M et al. (2009) The palliative prognostic score and survival in patients with advanced solid tumors receiving chemotherapy. Support Care Cancer 16:359-370
- [20] Kornblith PL. (2002) What is Possible When There is Evidence? One Linger in Palliative Medicine When Randomized Controlled Trials are Not Possible. Journal of Pain and Palliative Care Pharmacotherapy 23:44-50
- [21] Judet DA, Rosenberg B. (1993) Comparison of the Gompertz and Weibull functions as descriptives of human mortality distributions and the restrictions. Math Aging Dev 3:1-31
- [22] Szasz R, Szasz A. (2002) Biological implications of the Weibull and Gompertz models of aging. J Gerontol A Biol Sci Med Sci 57:569-578
- [23] Yuan YUN, Wei YI, Lin JI, Shih JI et al. (1995) Testing a model of aging: a animal experiments. Biometrika 82:353-372
- [24] Vansteenkiste JR, De Wever A, Braeckman BP. (1997) Two parametric logistic and Weibull equations provide better fits to survival data from biological populations. Journal of Clinical Oncology 15:3483-3490
- [25] Szasz R, Cole LD, Morgan D. (2003) Design on Phase II cancer trials evaluating survival probabilities. BMC Medical Research Methodology 3:18
- [26] Copas JW. De action of phase III trials and co-operative trials in randomised controlled trials of cancer: an analysis of a sigmoidal survival curve for muscle growth. Biometrika 70:177-197
- [27] Copas JW. A note on the effect of switching in hybrid designs suggest dependence on a phase transition. Physical Chem. and Phys. 12: 537-558
- [28] Copas JW. The limits of biological phase transition models as tested by survival curves: a review of approaches. Physical Chem. and Phys. 8: 579-597
- [29] Szasz R, E. J. Bartha and C. Cheung. 1996. The group sequential integral test for phase II cancer clinical trials. American Journal of Clinical Oncology 9:422-430